Mixed-Measurement Dynamic Factors: Systemic risk diagnostics, Coincident indicators and Early warning signals

Drew Creal[a], Bernd Schwaab[b]
Siem Jan Koopman[c,e], André Lucas[d,e]

[a]University of Chicago, Booth School of Business
[b]European Central Bank
[c]Department of Econometrics, Vrije Universiteit Amsterdam
[d]Department of Finance, Vrije Universiteit Amsterdam, and Duisenberg school of finance
[e]Tinbergen Institute, Amsterdam

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Motivation and contributions

- Economic and financial time series often share common features: business cycle dynamics.
- Time series can be continuous and/or discrete, and can be observed at different frequencies.
- We introduce observation-driven mixed measurement panel data models.
- We develop a model for macro-financial variables, credit ratings transitions and loss-given-default (LGDs).
- Dynamic factor model : simultaneous model with
  1. time-varying Gaussian model
  2. time-varying ordered logit
  3. time-varying beta distribution

Systemic risk indicators and early warning signals

- We aim to construct a coincident risk indicator and an early warning signal for financial distress
- Financial distress: exposure to simultaneous failure of many currently active financial intermediaries
- Our indicators are based on macro-financials, credit risk factors and loss-given-defaults (LGDs).
- Credit risk conditions can be decoupled from macro-financials
- Our credit risk factor has in the past preceded financial distress: early warning signal ?

Time varying parameter models

Consider autoregressive model : $y_t = \phi y_{t-1} + \varepsilon_t$ with $\varepsilon_t \sim \mathcal{N}(0, \sigma^2)$. Its prediction density is given by

$$y_t \ \sim \ p(y_t | Y^{t-1}; \psi), \qquad \psi = (\phi, \sigma^2)',$$

where $Y^t = \{y_1, \ldots, y_t\}$: basis for likelihood evaluation.

Replace σ^2 by a time-varying $\sigma_t^2 = f_t$, we then may consider the prediction density and time-varying specification

$$y_t \ \sim \ p(y_t | Y^{t-1}, F^{t-1}, f_t; \psi), \qquad \psi = \phi,$$

where $F^t = \{f_1, \ldots, f_t\} = \{F^{t-1}, f_t\}$.

time-varying model represented as a prediction density.

Time varying parameter models

Consider a sequence of observations $y_1, \ldots, y_T$ with density

$$y_t \quad \sim \quad p(y_t | Y^{t-1}, F^{t-1}, f_t; \psi),$$

which depends on a vector of time varying parameters f_t.

The parameters evolve through time as a linear process

$$f_{t+1} \quad = \quad \omega + B f_t + A s_t.$$

- $Y^t = \{y_1, \ldots, y_t\}$, $F^t = \{f_1, \ldots, f_t\}$.
- f_t contains all time varying parameters in observation density.
- ψ contains remaining fixed coefficients plus ω, A and B.
- s_t is the “driving” mechanism.
- exogenous variables (covariates) can also be included.

Time varying parameter models

Our general dynamic model is given by

$$\begin{aligned} y_t &\sim p(y_t | Y^{t-1}, F^{t-1}, f_t; \psi), \\ f_{t+1} &= \omega + Bf_t + As_t, \end{aligned}$$

Two options for dealing with s_t (Cox, 1981):

- let s_t be an i.i.d. random sequence $\Rightarrow$ parameter driven
- s_t is a deterministic function of Y^t $\Rightarrow$ observation driven

In case of an "observation driven" approach, the question is :

What is an appropriate function $g(\cdot)$ in $s_t = g(Y^t)$?

Time varying parameter models

In our work, we build observation driven models

$$
\begin{aligned}
y_t &\sim p(y_t|Y^{t-1}, F^{t-1}, f_t; \psi), \\
f_{t+1} &= \omega + Bf_t + As_t,
\end{aligned}
$$

where the evolution of the factors are determined by the scaled score

$$
\begin{aligned}
s_t &= S_t \cdot \nabla_t, \\
\nabla_t &= \frac{\partial \ln p(y_t|Y^{t-1}, F^{t-1}, f_t; \psi)}{\partial f_t}, \\
S_t &= S(t, Y^t; \psi).
\end{aligned}
$$

- motivation : update is similar to a Newton-Raphson step at time t.
- s_t is a m.d.s. which acts as a natural sequence of innovations.
- Creal, Koopman, Lucas (2008) propose this class of models and label them generalized autoregressive score (GAS) models.
- Different choices for S_t: Fisher info. matrix, identity matrix,....

Example: GARCH

Consider model $y_t \sim \mathcal{N}(0, \sigma_t^2)$:

$$y_t = \sigma_t \varepsilon_t \qquad \varepsilon_t \sim \mathcal{N}(0,1)$$

Let $f_t = \sigma_t^2$. The score and inverse of the information matrix are:

$$\begin{aligned} \nabla_t &= \frac{1}{2f_t^2} y_t^2 - \frac{1}{2f_t}, \\ S_t &= \mathcal{I}_{t-1}^{-1} = 2f_t^2. \end{aligned}$$

Here, the GAS$(1,1)$ recursion reduces to the GARCH$(1,1)$ model:

$$f_{t+1} = \omega + A_1(y_t^2 - f_t) + B_1 f_t$$

($A_1 = \alpha$ and $B_1 = \alpha + \beta$ from standard GARCH parameterization)

Special cases of this idea

- GARCH with normal dist.: Engle (1982), Bollerslev (1986)
- Exponential dist. (E-ACD and ACI): Engle & Russell (1998) and Russell (2001), respectively
- Gamma dist. (MEM): Engle (2002), Engle & Gallo (2006)
- Poisson dist.: Davis, Dunsmuir & Street (2003)
- Multinomial dist. (ACM): Russell & Engle (2005)
- Binomial dist.: Cox (1956), Rydberg & Shephard (2002)
- Student's t dist.: Harvey & Chakravarty (2008), Creal, Koopman & Lucas (2008, 2011)
- Beta and ordered logit dist.: this paper.

Mixed measurement panel data models

We introduce mixed measurement observation driven models

$$\begin{aligned} y_{it} &\sim p_i(y_{it}|Y^{t-1}, F^{t-1}, f_t; \psi), \quad i = 1, \ldots, N, \\ f_{t+1} &= \omega + B_1 f_t + A_1 s_t \end{aligned}$$

The score function is

$$\begin{aligned} s_t &= S_t \nabla_t \\ \nabla_t &= \sum_{i=1}^{N} \delta_{it} \nabla_{i,t} = \sum_{i=1}^{N} \delta_{it} \frac{\partial \log p_i(y_{it}|Y^{t-1}, F^{t-1}, f_t; \psi)}{\partial f_t}, \end{aligned}$$

- The observations y_{it} may come from different distributions.
- The factors f_t may be common across distributions.
- KEY: The score function allows us to pool information from different observations to estimate the common factor f_t.
- δ_{it} is an indicator function equal to 1 if y_{it} is observed and zero otherwise. Missing values are naturally taken into account.

Scaling matrix

Consider the eigenvalue-eigenvector decomposition of Fisher's (conditional) information matrix

$$\mathcal{I}_t \quad = \quad \mathrm{E}_{t-1}[\nabla_t \nabla_t'] \;=\; U_t \Sigma_t U_t',$$

The scaling matrix is then defined as

$$S_t \quad = \quad U_t \Sigma_t^{-1/2} U_t'$$

- S_t is then the "square root" of a generalized inverse.
- The innovations s_t driving f_t have an identity covariance matrix, when the info. matrix is non-singular.
- The conditional information matrix is additive for our models:

$$\mathcal{I}_t \quad = \quad \mathrm{E}_{t-1}[\nabla_t \nabla_t'] \;=\; \sum_{i=1}^{N} \delta_{it} \mathrm{E}_{i,t-1}[\nabla_{it} \nabla_{it}'].$$

Log-likehood function and ML estimation

- The log-likelihood function for an observation-driven model can easily be computed.
- The ML estimator is

$$\hat{\psi} = \arg\max_{\psi} \sum_{t=1}^{T} \sum_{i=1}^{N} \delta_{it} \log p_i(y_{it} | Y^{t-1}, F^{t-1}, f_t; \psi),$$

- Estimation is similar to a GARCH model.

Credit risk

- Growing econometrics literature on models for credit risk: McNeil et al. (2005), Bauwens and Hautsch (JFEct, 2006), Gagliardini and Gourieroux (JFEct, 2005), Koopman Lucas and Monteiro (JEct, 2008), Duffie et al. (JFE, JoF 2008).
- Basic observations:
  1. Probability of default varies over time with the business cycle.
  2. Conditional on default, the loss (recovery rate) varies with the business cycle.
  3. We observe excess clustering of defaults and ratings transitions beyond what can be explained by simply adding covariates.
  4. The literature focuses on a credit risk or frailty factor.
- Industry standard models are too simple to capture these features.
- New models in the literature are parameter driven models requiring simulation methods for estimation.
- We provide observation driven alternatives.

Data: Moody's and FRED

- We observe data from Jan. 1980 to March 2010.
- $7,505$ companies are rated by Moody's.
- We pool these into 5 ratings categories (IG, BB, B, C, D).
- We observe transitions, e.g. IG $\rightarrow$ BB or C $\rightarrow$ D
- There are $J = 16$ total types of transitions.
- 19,450 total credit rating transitions.
- 1,342 transitions are defaults.
- 1,125 measurements of loss-given default (LGD).
- LGD is the fraction of principal an investor loses when a firm defaults.
- We also observe six macroeconomic variables: industrial production growth, credit spread, unemployment, annual S&P500 returns, realized volatility, real GDP growth (qtly).

Models

- Credit ratings can be modeled using the (static) ordered probit model of CreditMetrics; one of the current industry standards, see Gupton Stein (2005).
- LGD's are often modeled by (static) beta distributions.
- GOAL: Build models that improve on current industry standards and are (relatively) easy to implement and estimate.
  1. Time-varying Gaussian
  2. Time-varying ordered logit
  3. Time-varying beta distribution
- Forecasting credit risk.
- Simulation of loss distributions and scenario analysis.
- Bank executives and regulators and can use them for "stress testing."

Mixed measurement model for credit risk

$$
\begin{aligned}
y_t^m &\sim \mathrm{N}\left(\mu_t, \Sigma_m\right) \\
y_{i,t}^c &\sim \text{Ordered Logit}\left(\pi_{ijt}, j \in \{\text{IG, BB, B, C, D}\}\right), \\
y_{k,t}^r &\sim \text{Beta}\left(a_{kt}, b_{kt}\right), \qquad k = 1, \ldots, K_t,
\end{aligned}
$$

- y_t^m are the macro variables.
- $y_{i,t}^c$ are indicator variables for each credit rating j for firm i.
- $y_{k,t}^r$ are the LGDs for the k-th default.
- K_t are the number of defaults in period t.
- μ_t, π_{ijt}, and (a_{kt}, b_{kt}) are functions of an $M \times 1$ vector of factors f_t.

Time varying Gaussian model for macro data

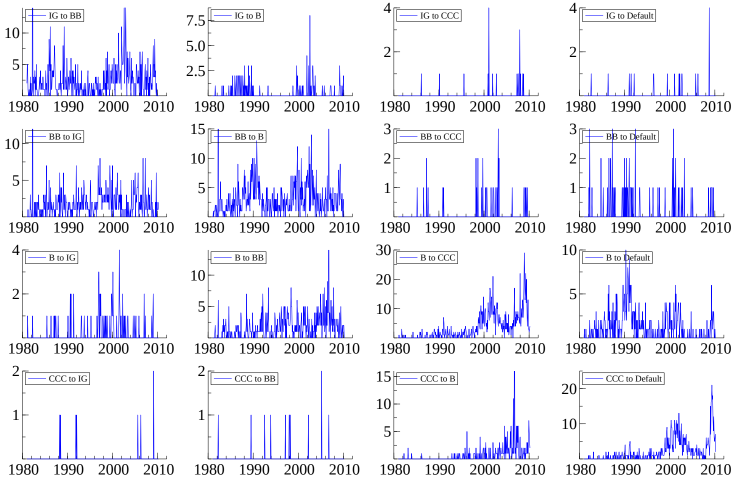
$$
\begin{aligned}
y_t^m &\sim \mathrm{N}\left(\mu_t, \Sigma_m\right), \\
\mu_t &= Z^m f_t.
\end{aligned}
$$

- Z^m is a $(6 \times M)$ matrix of factor loadings.
- Σ_m is a (6×6) diagonal covariance matrix.
- $\tilde{S}_t$ is a selection matrix indicating which macro variables are observed at time t.

$$
\begin{aligned}
\nabla_t^m &= \left(\tilde{S}_t Z^m\right)' \left(\tilde{S}_t \Sigma_m \tilde{S}_t'\right)^{-1} \tilde{S}_t \left(y_t^m - \mu_t\right), \\
\mathcal{I}_t^m &= \left(\tilde{S}_t Z^m\right)' \left(\tilde{S}_t \Sigma_m \tilde{S}_t'\right)^{-1} \tilde{S}_t Z^m.
\end{aligned}
$$

Moody's monthly credit ratings transitions

The data have been pooled together each month.



Time-varying ordered logit

$$
\begin{aligned}
y^c_{i,t} &\sim \text{Ordered Logit}\left(\pi_{ijt}, j \in \{\text{IG, BB, B, C, D}\}\right), \\
\pi_{ijt} &= \mathrm{P}\left[R_{i,t+1} = j\right] = \tilde{\pi}_{ijt} - \tilde{\pi}_{i,j-1,t}, \\
\tilde{\pi}_{ijt} &= \mathrm{P}\left[R_{i,t+1} \leq j\right] = \frac{\exp(\theta_{ijt})}{1+\exp(\theta_{ijt})}, \\
\theta_{ijt} &= z^c_{ijt} - Z^{c\prime}_{it} f_t.
\end{aligned}
$$

- $J^c = 5$ categories $j \in \{\text{IG, BB, B, C, D}\}$.
- R_{it} is the rating for firm i at the start of month t.
- y^c_{it} is an indicator variable for each rating type.
- π_{ijt} is the probability that firm i is in category j.
- $\tilde{\pi}_{i,\text{D},t} = 0$ and $\tilde{\pi}_{i,\text{IG},t} = 1$.
- To our knowledge, a time-varying ordered logit model is new.

Time-varying ordered logit

The contribution to the log-likelihood at time t is

$$\ln p_i(y_{it}^c | F^t, Y^{t-1}; \psi) \quad = \quad \sum_{i=1}^{N_t} \sum_{j=1}^{J^c} y_{ijt}^c \log\left(\pi_{ijt}\right)$$

The score and information matrices are

$$\begin{aligned}
\nabla_t^c \quad &= \quad -\sum_{i=1}^{N_t} \sum_{j=1}^{J^c} \frac{y_{ijt}^c}{\pi_{ijt}} \cdot \dot{\pi}_{ijt} \cdot Z_{it}^c, \\
\mathcal{I}_t^c \quad &= \quad \sum_{i=1}^{N_t} n_{it} \left(\sum_j \frac{\dot{\pi}_{ij,t}^2}{\pi_{ij,t}} \right) Z_{it}^c Z_{it}^{c\prime}
\end{aligned}$$

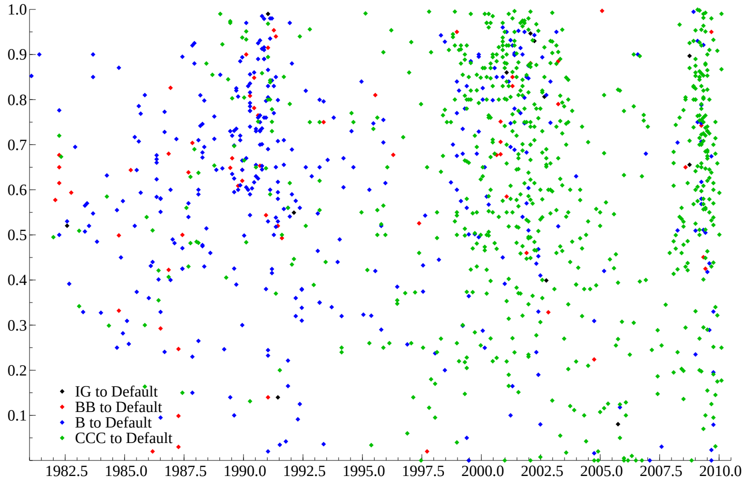
where

$$\dot{\pi}_{ijt} = \tilde{\pi}_{ijt}\left(1 - \tilde{\pi}_{ijt}\right) - \tilde{\pi}_{i,j-1,t}\left(1 - \tilde{\pi}_{i,j-1,t}\right).$$

Loss given default

- When a firm defaults, investors typically lose a fraction of their investment (alternatively, they recover a fraction of their investment).
- The fraction of losses experienced by investors also varies with the business cycle.
- We develop a new model for a time-varying beta distribution.
- See McNeil and Wendin (2007 JEmpFin) for Bayesian inference in a state space model.

Loss given default by transition type



Time-varying beta distribution

$$
\begin{aligned}
y^r_{k,t} &\sim \mathrm{Beta}\left(a_{kt}, b_{kt}\right), \qquad k = 1, \ldots, K_t, \\
\\
a_{kt} &= \beta_r \cdot \mu^r_{kt} \\
b_{kt} &= \beta_r \cdot (1 - \mu^r_{kt}) \\
\\
\log\left(\mu^r_{kt} / \left(1 - \mu^r_{kt}\right)\right) &= z^r + Z^r f_t.
\end{aligned}
$$

- We observe $K_t \geq 0$ defaults at time t.
- $0 < y^r_{k,t} < 1$ is the amount lost conditional on the k-th default.
- μ^r_{kt} is the mean of the beta distribution.
- z^r is the unconditional level of LGDs.
- Z^r is a $(1 \times M)$ vector of factor loadings.
- β_r is a scalar parameter

Time-varying beta distribution

The contribution to the log-likelihood at time t is

$$\begin{aligned}\ln p_i(y^r_{kt}|F^t, Y^{t-1};\psi) &= \sum_{k=1}^{K_t}(a_{kt}-1)\log\left(y^r_{kt}\right)+(b_{kt}-1)\log\left(1-y^r_{kt}\right)\\ &\quad -\log\left[B\left(a_{kt},b_{kt}\right)\right]\end{aligned}$$

The score and information matrices are

$$\begin{aligned}\nabla^r_t &= \beta_r\sum_{k=1}^{K_t}\mu^r_{kt}(1-\mu^r_{kt})\left(Z^r\right)'(1,-1)\left((\log(y^r_{kt}),\log(1-y^r_{kt}))'-\dot{B}\left(a_{kt},b_{kt}\right)\right)\\ \mathcal{I}^r_t &= \beta_r\sum_{k=1}^{K_t}(\mu^r_{kt}(1-\mu^r_{kt}))^2\left(Z^r\right)'(1,-1)\left(\ddot{B}\left(a_{kt},b_{kt}\right)\right)(1,-1)'Z^r\end{aligned}$$

where

$$\sigma^2_{kt}=\mu^r_{kt}\cdot(1-\mu^r_{kt})/(1+\beta_r).$$

Estimation details

- The macro data y_t^m has been standardized.
- We consider models with $p = 1$ and $q = 1$ factor dynamics.
- For identification of the level parameters, we set $\omega = 0$ in the factor recursion:

$$f_{t+1} = A_1 s_t + B_1 f_t$$

- For identification of the factors, we also impose restrictions on Z^m, Z^c, and Z^r.
- Some parameters have been pooled for "rare" transitions; e.g., IG $\rightarrow$ D and BB $\rightarrow$ D.
- Moody's re-defined several categories in April 1982 and Oct. 1999 causing incidental re-ratings (outliers), which we handle via dummy variables for these dates.

AIC, BIC, and log-likelihoods for different models

	(2,0,0)	(2,1,0)	(2,2,0)	(3,0,0)
log-Like	-40447.9	-40199.1	-40162.8	-40056.2
AIC	81005.9	80520.1	80457.0	80242.4
BIC	81640.0	81223.0	81218.0	80991.0
	(3,1,0)	(3,2,0)	(3,1,1)	(3,2,1)
log-Like	-39817.1	-39780.8	-39812.6	**-39780.0**
AIC	79776.2	**79713.6**	79771.2	79716.0
BIC	80594.0	**80589.0**	80612.0	80615.0

The number of factors for each data type are represented by (m, c, r).

Parameter estimates for the (3,2,0) model

Macro loadings Z^m

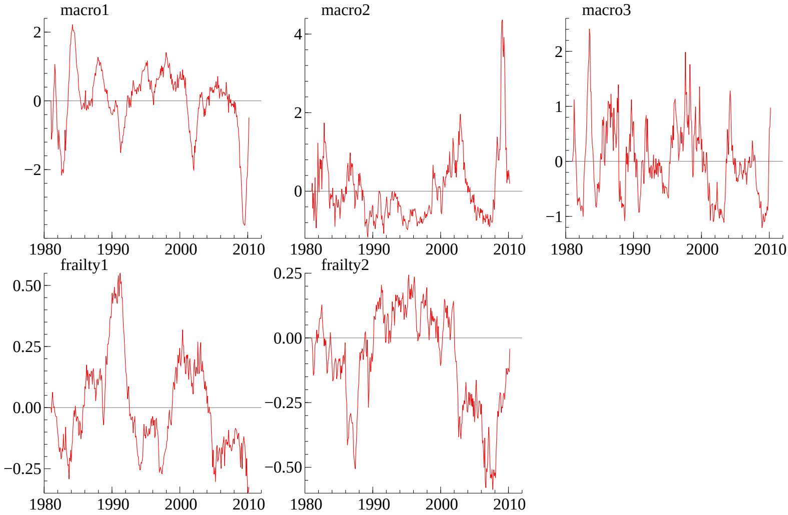
	macro$_1$	macro$_2$	macro$_3$	frailty$_1$	frailty$_2$
IP	1.000	0.000	0.000	0.000	0.000
	——	——	——	——	——
UR	-0.892***	0.122***	-0.062*	0.000	0.000
	(0.037)	(0.041)	(0.040)	——	——
RGDP	0.811***	0.072	0.336***	0.000	0.000
	(0.066)	(0.079)	(0.074)	——	——
Cr.Spr.	-0.169**	1.000	0.000	0.000	0.000
	(0.085)	——	——	——	——
$r_{S\&P}$	0.049	-0.268***	1.223***	0.000	0.000
	(0.093)	(0.081)	(0.093)	——	——
$\sigma_{S\&P}$	-0.007	0.648***	1.000	0.000	0.000
	(0.107)	(0.084)	——	——	——

Parameter estimates for the (3,2,0) model

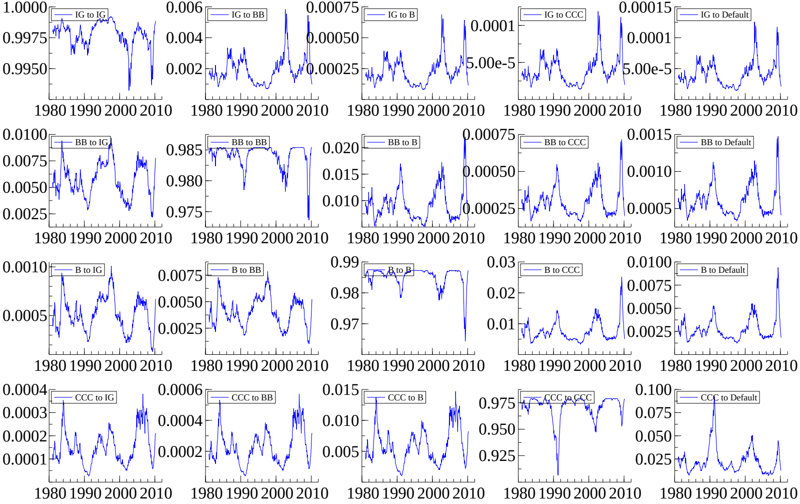
Credit rating and LGD loadings Z^c and Z^r

	macro_1	macro_2	macro_3	frailty_1	frailty_2
Z^c					
IG	-0.052	0.202***	-0.123**	1.475***	-1.165**
	(0.059)	(0.055)	(0.069)	(0.371)	(0.555)
BB	-0.078**	0.172***	-0.102***	1.000	0.000
	(0.037)	(0.037)	(0.040)	—	—
B	-0.184***	0.162***	-0.142***	0.970***	-0.016
	(0.035)	(0.031)	(0.040)	(0.156)	(0.158)
CCC	-0.262***	0.073*	-0.018	1.936***	1.000
	(0.057)	(0.050)	(0.075)	(0.465)	—
Z^r	0.018	0.276***	-0.082*	1.212***	1.065***
	(0.049)	(0.046)	(0.062)	(0.376)	(0.301)

Estimated factors for the (3,2,0) model



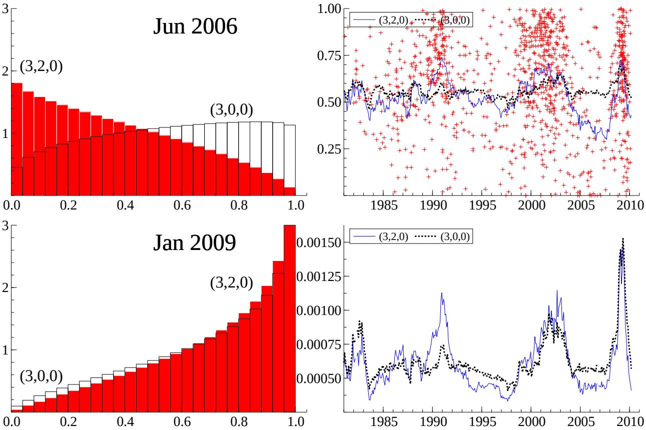
Time-varying transition probabilities



Simulating cumulative loss distributions

- Most financial institutions carry a large portfolio of credit related securities.
- Given a portfolio at time T, we can use the models to simulate different possible risk scenarios.
- GOAL: determine the amount of capital banks may need in the future.
- What happens if we do not include time-varying parameters f_t in the model?
- Scenario analysis:
  1. What happens if there is a negative shock to RGDP?
  2. What happens if there is an increase to credit spreads?

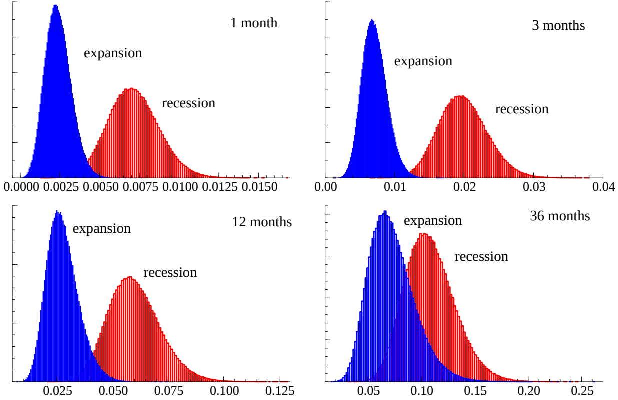
Loss given default results



Top and bottom left are loss distributions. Top right is a plot of the mean through time. Bottom right are transition probabilities from BB $\rightarrow$ D.

Simulating cumulative loss distributions

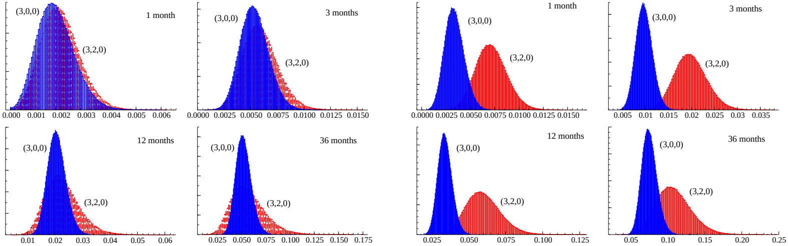
Cumulative losses on a portfolio of bonds at different horizons.



Comparison between a recession and expansion for (3,2,0) model.

Simulating cumulative loss distributions

Comparison of cumulative loss distributions with/without factors.



Left: starting at $f_T = 0$. Right: starting in a recession.

Conclusion and future work

- We introduce a new class of observation-driven models for mixed-measurement data which share exposure to common factors.
- Missing values and mixed frequencies are handled in a natural way.
- Using this approach, we develop new models for credit risk.
- The models can be used for simulating loss distributions, stress testing, and scenario analysis.
- Future work:
  - When computing loss distributions, current models do not account for changes in market prices of bonds or loans.
  - Current models depend on industry credit ratings by Moody's, Fitch, Standard & Poors.
  - Potential to use alternative sources of data.